

## WORKING PAPERS

Macro Determinants of Individual Income Poverty in 93 Regions of Europe

> Anne REINSTADLER<sup>1</sup> Jean-Claude RAY<sup>2</sup>

CEPS/INSTEAD, Luxembourg<sup>1</sup> Nancy University and CNRS UMR, France<sup>2</sup>

CEPS/INSTEAD Working Papers are intended to make research findings available and stimulate comments and discussion. They have been approved for circulation but are to be considered preliminary. They have not been edited and have not been subject to any peer review.

The views expressed in this paper are those of the author(s) and do not necessarily reflect views of CEPS/INSTEAD. Errors and omissions are the sole responsibility of the author(s).

## Macro Determinants of Individual Income Poverty in 93 regions of Europe

Anne REINSTADLER<sup>1</sup> CEPS/INSTEAD, Luxembourg

Jean-Claude RAY<sup>2</sup> Nancy University and CNRS UMR 7522, France

*Abstract*: The analysis of the at-risk-of-poverty determinants can be improved by taking into account factors at macro (regional) level. This hypothesis has already been made in previous research, at country-level, on cross-sectional data. We use longitudinal data in this analysis in order to get more precise estimated parameters, and we test if the regional unemployment rate and the regional GDP affect the individual at-risk-of-poverty status. The countries taken into account are those present in the Statistics on Income and Living Conditions (EU-SILC) dataset.

Key words: income poverty, EU-SILC, multilevel models, longitudinal data

*JEL*: I32

Avril 2010

<sup>&</sup>lt;sup>1</sup> Corresponding author: anne.reinstadler@ceps.lu

<sup>&</sup>lt;sup>2</sup> jean-claude.ray@univ-nancy2.fr

We would like to thank David Brady, Jacques Brosius, Alessio Fusco, Tony Atkinson, Eric Marlier and Philippe Van Kerm for useful discussions led during this project. This paper has been presented at the Net-Silc International Conference, 25-26 March 2010, Warsaw. Of course, remaining errors are ours.

#### 1. Introduction

Tackling poverty by 2010 was one of the European objectives defined by the Lisbon European Council in 2000. Ten years later, 2010 is the European year for combating poverty and social exclusion. Poverty continues therefore to be at the heart of social policy in most European Member States. Ideally, social policies aimed at reducing poverty need to be based on an in-depth understanding of the underlying processes at work. A first step towards such an understanding consists in shedding some light on the main determinants of poverty.

Early descriptive studies have checked for relationships between poverty status and different characteristics taken in turn (Bradshaw, 1999; Bradubury et al., 1999; UNICEF, 2000; Mejer et al., 2000). This has given some insight on the factors involved, but these studies have only partially allowed to understand how these factors work. Other researches (see for example Cappellari and Jenkins, 2002; Fertig and Tamm, 2007; Brady et al., 2009) have extended this initial approach by reasoning all other things being equal, checking the effect on poverty of factors such as educational attainment, age, employment status, family structure – all of these factors having been calculated at the individual<sup>3</sup> level. Simultaneously another stream of studies (see Moller et al., 2003; Wiepking and Maas, 2005; Brady et al., 2009; Tai and Treas, 2008) has emphasized the analysis of the role of macro characteristics in a cross-national context. These analyses have shown that the macro factors could well have an effect on the poverty probability. Indeed, the generosity of social benefits (and especially of family benefits) proves to have a significant negative effect on the odds of poverty (see Brady et al. 2009, Moller et al. 2003).

For all of these three types of analyses, one major improvement has consisted in taking into account the longitudinal feature of poverty, using panel data<sup>4</sup>. The indicator of persistent poverty<sup>5</sup>, for example, allows to figure out whether poverty is a temporary or rather a long-term phenomenon. Furthermore, developments in the econometrics of panel data have allowed researchers to further investigate important topics such as poverty duration or unobserved heterogeneity.

However, to our knowledge, no study has yet dealt simultaneously with all EUcountries, longitudinal data and factors at both individual and macro levels using a relevant specification. Brady et al. (2009) study the effect of macro-determinants on the probability of being poor using a GEE model<sup>6</sup> applied to 15 EU (plus some non-EU) countries but using cross-sectional data. In this paper we will extend this kind of analysis to 93 EU-regions (26 countries) and, contrary to previous work, we will use a longitudinal dataset (EU-SILC 2005 and 2006).

This paper has the following objective: it aims at disentangling the role of micro and macro factors in explaining the poverty status, by using detailed information about different

 $<sup>^{3}</sup>$  As the same poverty status is, by the European definition, affected to all individuals belonging to the same household, some authors have defined the factors exclusively at that level (see for example Andriopoulou et al., 2008).

<sup>&</sup>lt;sup>4</sup> See Ray and Jeandidier (2003) for a comprehensive review of the French literature on this subject.

<sup>&</sup>lt;sup>5</sup> This indicator belongs to the set of common indicators for the social protection and social inclusion process adopted by the Social Protection Committee in 2006.

<sup>&</sup>lt;sup>6</sup> A Generalized Estimated Equations model can be used to estimate marginal or population-averaged effects taking into account the dependence among units nested in clusters (Rabe-Hesketh and Skrondal, 2008).

regions in Europe<sup>7</sup>. Indeed, we would like to test if there is a genuine effect of macro factors such as the unemployment rate on the poverty probability, and especially if these factors can affect the impact of individual characteristics such as the education level on this probability.

In Section 2 we present the definition of income poverty that we will apply. In Section 3 we then develop the different econometric methods available to deal with the question and data at hand. Section 4 gives a detailed description of our dataset. The results and comments of our own model are then presented in Section 5. Final conclusions are to be found in Section 6.

#### 2. The definition of income poverty in Europe

In Europe, poverty is officially defined in relative terms, as the percentage of individuals living in a household whose equivalent income is below the poverty threshold. This threshold is defined in each country (equal to 60% of the national median equivalent income), aiming at taking into account the national income inequalities. As a consequence, two countries with very different standards of living (and thus very different median equivalent income and different poverty thresholds) can have the same poverty rate.

Seemingly contradictory results due to this definition do not matter as long as one is aware of the conventions they are based on, and when the at-risk-of-poverty rates are interpreted together with the threshold values. But, in our case, the main objective is to figure out to what extent the poverty status is explained by some macro factors such as the unemployment rate. It is thus necessary, in order to allow that kind of relationship to appear, that the poverty indicator ranks the countries as the macro variables do. As a consequence, we have chosen to keep defining poverty in a relative way (60% of a certain threshold) but to calculate a new threshold, allowing this kind of ranking. With this objective in mind, we calculate a unique European poverty threshold by considering all individuals to belong to a same big country, which is Europe<sup>8</sup>.

By doing this, we move away from the official EU-definition of the at-risk-of-poverty status (which is rather a measure of intra-country inequality), and we consider Europe as an integrated entity. In the same spirit, we exclude neither Iceland nor Norway: while not part of the EU-27 in 2009, they could be expected to join.

<sup>&</sup>lt;sup>7</sup> Indeed it is between regions, and not between individuals, that some variance of macro factors can be found. This is probably also true between countries but because our data is limited with respect to the number of available countries, we prefer the analysis at the regional level.

<sup>&</sup>lt;sup>8</sup> Using the usual at-risk-of-poverty threshold (which gives similar at-risk-of-poverty rates for countries with very different economic situations) could lead to this kind of situation: the characteristics of the countries would not be relevant to explain their own at-risk-of-poverty rate because the latter is more a measure of income inequalities than a measure of economic performance. In other words, macro economic factors, which can be considered from a theoretical point of view to be associated with poverty, can not be linked with the individual poverty status such as defined when using the official definition. Note that this definition of a European at-risk-of-poverty threshold is supported by Marlier et al. (2007).

We then determine, for each country, the fraction of individuals<sup>9</sup> situated below this new European threshold (see Appendix 1, where these figures can be compared with the official at-risk-of-poverty rates, based on the national thresholds).

#### 3. Methodology

Two major approaches have been used to study the determinants of poverty. One consists in explaining the transitions into and out of poverty (probability of staying poor, and probability of entering poverty). The second approach focuses on the poverty status at a specific point in time.

The first approach takes into account the initial conditions problem<sup>10</sup> by using longitudinal data. This problem refers to the fact that the poverty status during the first period may not be exogenous because of observed and unobserved characteristics, which would affect the probability of being poor afterwards. However, papers running that kind of analysis do not introduce macro factors in the analysis. This is either because they are interested in only one country (see Cappellari and Jenkins, 2002, 2004; Van Kerm, 2004; Buddelmeyer and Verick, 2007; Ayllon, 2008<sup>11</sup>), or because the different countries are treated separately, with as many models as there are countries (Andriopoulou et al., 2008).

On the other hand, some authors estimate the probability of being poor at a specific point in time (see Wiepking and Maas, 2005; Brady et al., 2009; Tai and Treas, 2008). All of these authors use cross-sectional data of 22 countries from the Luxembourg Income Study (LIS) and integrate macro variables in the analysis (such as the unemployment rate or the welfare generosity), stressing that the welfare system could play a role in allowing individuals to escape from  $poverty^{12}$ .

As for our own work, it focuses on the poverty status and integrates macro factors as well. Our method is original in two ways: first, it uses longitudinal data and second, it takes into account the fact that some variability can be found at the regional level. In fact, two reasons have led us to do the analysis at the regional rather than at the country-level: first because the situation the individuals face (in terms of unemployment rate for example) could be very different from one region to another, within the same country, and second because there are more regions than countries (93 versus 26), which is better from a statistical point of view<sup>13</sup>.

<sup>&</sup>lt;sup>9</sup> These figures concern working age individuals (aged 25-55) because the behaviours and thus the factors at work can be very different for both other groups (children and retired people).

<sup>&</sup>lt;sup>10</sup> Some authors also control for the retention probability (see for example Cappellari, 2002; Cappellari and Jenkins, 2002, 2004; Ayllon, 2008). The idea is that the probability of being observed during two consecutive years could depend on unobserved characteristics of the individuals that should thus be controlled for. <sup>11</sup> Other authors have used the same kind of models, on related subjects but not poverty: Stewart and Swaffield

<sup>(1999)</sup> and Cappellari (2004) on earning, Poggi (2007) on social exclusion persistence.

<sup>&</sup>lt;sup>12</sup> Moller et al. (2003) work on these data as well, but at the macro level. Indeed, they link the national at-riskof-poverty rate to macro variables such as GDP or the employment rate in the agricultural sector. Their sample is guite small (61 observations, nested in 14 countries).

<sup>&</sup>lt;sup>13</sup> While having 93 higher-level groups is technically much better than using only 26 higher-level groups, especially as far as variances/covariances estimation is concerned, a problem remains here: regions are nested within countries. This suggests to move from a three-level analysis (observations over time nested within individuals, themselves nested within regions) to a four-level analysis (adding the country-level at the top of the

In other words, we estimate a model of poverty probability, using two years of observations for each individual (in order to increase the estimation accuracy). Some of these individuals live in the same region<sup>14</sup>. From an econometric point of view, this data setup leads to a problem concerning the independence of observations: individuals being observed over two years and/or living in the same region share their own time-invariant characteristics and/or the characteristics of the area and can therefore no longer be considered to be independent. As a consequence, using traditional techniques would give consistent estimates but heavily<sup>15</sup> under-estimated standard errors.

In order to cope with that statistical problem, we have chosen one of the many available techniques: we run a multilevel model, which treats the upper levels (the individuals and the regions here) not as unique entities but as units primarily characterised by factors calculated at their level (e.g. characteristics of the individual, or the unemployment rate of the region). These models explicitly take into account the hierarchical structure of the data, thereby allowing us to analyse — first to measure, then to explain — the fraction of the variability of the poverty rate which is attached to each (nested) level. Contrary to the fixed effects models<sup>16</sup>, multilevel models make use of the between variance, and are therefore especially useful when this variance is quite high. Some authors have already stressed that the use of this kind of models would be relevant in this framework (Cappellari and Jenkins 2004; Brady et al. 2009) but they have underlined the complexity of these models, whose convergence status is often out of reach.

The model we estimate is a binary logistic regression, where the probability of being at-risk-of-poverty is explained. This multilevel model takes into account three levels: time (measured in years), individuals and regions. It can be written as follows in its structural form  $^{17}$ :

 $logit(P_{ijk}) = \beta_{0k} + \beta_{1k} x_{1ijk} + \beta_2 x_{2ijk}$  $\beta_{0k} = \gamma_0 + \gamma_1 z_{0k} + U_{0k}$  $\beta_{1k} = \delta_0 + \delta_1 z_{1k} + U_{1k}$ 

hierarchy). But models with four levels and so many observations prove to have convergence problems and can thus not be estimated. We have therefore chosen to restrict our analysis to three levels, and to consider the third as being either the regions (in which case the country variable is introduced in the model as an explanatory variable) or the countries (the regional level being then put aside – see Appendix 2).

<sup>&</sup>lt;sup>14</sup> Some of them also live in the same household. Ideally, this fact should be taken into account as well, but again it would increase the complexity of the analysis: we would have to deal with 3 sources of non independence of the observations (individuals observed several years, living in the same household, and in the same region). In this paper, we cope with two of them.

<sup>&</sup>lt;sup>15</sup> According to what Angrist and Pischke (2009) name "the Moulton factor" (p. 310), we calculate to what extent the macro factors regression coefficients would be overestimated by ignoring intraregional correlation. We use the general formula 8.2.5 (ibidem p. 311) which allows for various cluster sizes. With an intra-class correlation coefficient (given by the empty model – see Appendix 6) of 0.551, a (non weighted) average regional sample size of 3206.20 individuals, a (non weighted) variance of regional sample size amounting to 15,550,017.33 and an intra-class correlation of macro factors equal to 1 by definition, we get 66.63. Note that this very large impact of clustering on standard errors is here due to the conjunction of big discrepancies between the size of regions (from 19 to 19941) and a high intra-class correlation coefficient (0.551). Even if the Moulton formula should be considered here only as a very rough approximation (because this formula ignores weighting and repeated observations), it clearly suggests that clustering effects must be taken into account.

<sup>&</sup>lt;sup>16</sup> Yet these models have a strong advantage: they control for group-invariant factors, measured or unmeasured. But this advantage has a price: the inability to estimate regression coefficients for these group-invariant factors, and thus to allow the analyst to conclude in terms of the effect of these factors.

<sup>&</sup>lt;sup>17</sup> The notation we use here is the same as Snijders and Bosker's (2004).

where:

i indexes time

j indexes the individuals

k indexes the regions

 $P_{ijk}$  is the probability of being at-risk-of-poverty

 $x_{1ijk}$  is a vector of independent factors defined at the individual/household level whose effects are assumed to be random

 $x_{2ijk}$  is a vector of independent factors defined at the individual/household level whose effects are assumed to be fixed

 $z_{0k}$  is a vector of independent factors defined at the regional level, which are supposed to have an impact on the average  $P_{ijk}$  in region k

 $z_{1k}$  is a vector of independent factors defined at the regional level, which are supposed to moderate the effect of the  $x_{1ijk}$  on  $P_{ijk}$ 

 $\beta_{0k}$  is a random intercept

 $\beta_{1k}$  is a vector of random slopes

 $\beta_2$  is a vector of fixed slopes

 $\gamma_0$  measures the average value of  $P_{ijk}$  across regions, when each independent variable is 0

 $\gamma_1$  measures the impact of  $z_{0k}$  on  $P_{ijk}$ 

 $\delta_0$  measures the average impact, across regions, of  $x_{1ijk}$  on  $P_{ijk}$ , when each  $z_{1k}$  is 0

 $\delta_1$  measures the impact of  $z_{1k}$  on the effect of the  $x_{1ijk}$  on  $P_{ijk}$ 

 $U_{0k}$  and  $U_{1k}$  are error terms assumed to follow a multinormal distribution  $\mathbb{N}(0,0;\Omega)$ ,  $\Omega$  being the variance-covariance matrix<sup>18</sup>.

The reduced form is thus:

$$logit(P_{ijk}) = \gamma_0 + \gamma_1 z_{0k} + \delta_0 x_{1ijk} + \delta_1 z_{1k} x_{1ijk} + \beta_2 x_{2ijk} + U_{0k} + U_{1k} x_{1ijk}$$

This formula refers to a random slope model, meaning that the intercept and at least one of the explanatory variables have a random coefficient.

#### 4. Data

The EU-SILC longitudinal dataset provides information at both individual and household levels and covers at most 5 years (from 2003 to 2007) depending on the country: data are not available for some countries in 2003, 2004 and 2007. Had we used all five waves to calculate the poverty threshold as we define it (i.e. at the European level), it would have increased or decreased over time just because some countries (e.g. Germany) are absent for some years – it means, without any link with the economic situation. As a consequence, we work with data from two waves only, 2005 and 2006, where all 26 countries are present.

The unit of analysis is the individual: as stated by an OECD report (2001), this is the usual choice for poverty analysis with longitudinal data because individuals can be followed

<sup>&</sup>lt;sup>18</sup> In our main model, we specify an unstructured form of the variance-covariance matrix (allowing covariances between random effects to be non zero) because the covariance between the error terms of the intercept and the number of employed people in the household appears to be highly significant.

over time whereas households  $cannot^{19}$ . The sample contains 131 891 working age adults (25-55) for the first wave, and 166 379 for the second<sup>20</sup>, split between 26 countries (see Appendix 3). These countries<sup>21</sup> are in turn divided into 93 regions (see Appendix 4).

The explanatory variables have been chosen in order to control for different determinants of the poverty status. Some are related to the demographic characteristics of the household <sup>22</sup> (number of children and number of adults), others to the labour market (presence of at least one adult with an upper level of education, number of employed people), others still to the health status (presence<sup>23</sup> of at least one adult with chronic disease, or hampered by illness in his/her daily activities). Two additional variables are measured at the regional level: the GDP and the unemployment rate<sup>24</sup>.

Some descriptive statistics, for both the whole sample and each country, can be found in Appendix 5 (but not for each region: as there are 93 regions<sup>25</sup>, it proves not to be sensible to show the descriptive statistics for each of them).

Beyond the usual hypotheses concerning all the control variables<sup>26</sup>, we make two further ones concerning our variables of interest: first we assume that the negative effect of the level of education on the probability of being at-risk-of-poverty could be weaker in richer areas (where the probability of being poor is quite low, whatever the level of education). Second we assume that the negative effect of the number of employed people in the

<sup>&</sup>lt;sup>19</sup> As usual, the poverty threshold is calculated at the individual level but the poverty status is defined at the household level.

 $<sup>^{20}</sup>$  Note that the sample is not balanced: 41 % of the individuals have only one observation (only 5% in Denmark, up to 59% in Czech Republic).

<sup>&</sup>lt;sup>21</sup> The latest release of SILC longitudinal data (August 2009) contains only 22 countries, Germany, Ireland, Greece and Denmark being absent. But we wanted both to work on this release, the data of which have been cleared of previous problems, and to keep these four countries in the analysis. We have therefore added to these 22 countries the other 4 from the previous release (March 2009).

 <sup>&</sup>lt;sup>22</sup> All of the explanatory variables, apart from gender and age, are calculated at the household level, because the poverty status is defined at that level.
 <sup>23</sup> Several variables have originally been defined at the individual level (e.g. having a chronic disease). In order

<sup>&</sup>lt;sup>23</sup> Several variables have originally been defined at the individual level (e.g. having a chronic disease). In order to have all variables defined at the same (household) level, we have tried to build aggregated variables such as the number of adults in the household suffering from a chronic disease. Unfortunately, this information was available only for one individual per household in the register countries (which use information from administrative datasets when available and interview only one individual per household for the remaining questions to be asked). Keeping this kind of definition would thus have led to a big loss of information. As a consequence, we have defined a much less precise indicator, such as "presence in the household of at least one adult suffering from a chronic disease". Note that even this imprecise indicator could be ill measured in those register countries, as only one household member is interviewed (and the construction of the variable would then only rest on that member). The level of education is defined for all household members aged 16 or more, but the variable contains a lot of missing values in some countries (13% in Portugal, 14% in Spain and up to 16% in the UK in 2005). We have thus adopted the same definition in order to construct a variable at household level.

<sup>&</sup>lt;sup>24</sup> Unfortunately, we were not able to integrate in our main model potential important macro factors such as the social expenditures (expressed as percentages of the GDP): unemployment compensation, public health expenditures, expenditures with respect to inclusion. Indeed, they are not yet known at regional level.

 $<sup>^{25}</sup>$  In fact, 16 countries out of the 26 countries at hand only have one region, either because they are quite small (12 of them), or because the region code is not available in the dataset (4 of them).

<sup>&</sup>lt;sup>26</sup> Note that income (and thus the poverty status) and individual/household demographic characteristics have not been measured at the same time: income refers to the year prior to the survey, whereas demographic characteristics to the time of the survey. We could have dealt with this by lagging all non-income variables (e.g. demographic characteristics given for year 2006 should be linked to the income declared in 2007), yet at the price of losing some countries since not all of them have available data for 2007.

household on the probability of being at risk of poverty could be attenuated when the unemployment rate is high due to the downward pressure on wages.

#### 5. Results and Comments

The results of the model are shown in Table 1. All our analyses (descriptive and econometric) use weighted data  $^{27}$ .

<sup>&</sup>lt;sup>27</sup> Moon and Stotsky (1993) state that, if the data come from a stratified and clustered random sampling, it is reasonable to treat the sample as a simple random sample, thus ignoring weights. And Poggi (2007) states that it is more efficient, from an econometrical point of view, not to weigh the data. That way of doing has been adopted by studies on panel data (as stated by Ayllon, 2008, or Andriopoulou et al., 2008). But the question of weighting the data is still open). We have chosen to weigh them, as the SILC dataset is known not to be representative of the population.

			20
Table 1 Probability	v of being at-risk-of-ne	overty in 93 European region	s. Estimation with a multilevel model <sup>28</sup>
Table 1. Trobabilit	y of being at-fisk-of-pe	overty in 75 European region	3. Estimation with a multilevel model

Table 1. Probability of being at-risk-of-poverty in 93 European regions. Estimation with a multilevel model <sup>28</sup>						
	Parameter estimate	Standard error	Odds Ratios			
Variables <sup>29</sup>						
intercept	0,2625	0,3806				
woman	0,01578	0,0139	1,016			
age centered (around the average : 41.27)	0,002428 *	0,000848	1,002			
age centered squared	0,000453 ***	0,000105	1,000			
chronic disease in the household	-0,1214 ***	0,01692	0,886			
activity hampered by disease in the household	0,1612 ***	0,01754	1,175			
upper education level in the household	-1,5241 ***	0,1568	0,218			
number of children in the household	0,3555 ***	0,01532	1,427			
number of children in the household squared	-0,0127 *	0,004298	0,987			
number of adults in the household	-0,2607 ***	0,02786	0,770			
number of adults in the household squared	0,08983 ***	0,004659	1,094			
number of employed people in the household	-1,2964 ***	0,07399	0,274			
regional annual GDP per capita (in 10 <sup>3</sup> Euros)	-0,04818 ***	0,007353	0,953			
regional unemployment rate (expressed in %)	0,02103	0,01122	1,021			
upper education level * regional GDP per capita	0,02427 **	0,006689				
number of employed people * regional unemployment rate	0,01392 *	0,005996				
wave 2005	-0,2055 ***	0,01488	0,814			
wave 2006	ref.	ref.	ref.			
country BE	-0,3208	0,3616	0,726			
country CZ	1,9401 ***	0,3278	6,960			
country DK	-0,6816	0,4419	0,506			
country DE	-0,2740	0,4252	0,760			
country EE	3,1916 ***	0,4460	24,328			
country IE	-0,2453	0,4468	0,782			
country EL	0,9511 *	0,3431	2,589			
country ES	0,9605 *	0,3115	2,613			
country FR	-0,2864	0,3107	0,751			
country IT	0,2918	0,3311	1,339			
country CY	-0,2265	0,4889	0,797			
country LV	3,5168 ***	0,4476	33,678			
country LT	3,7270 ***	0,4445	41,556			
country LU	-0,3031	0,7205	0,739			
country HU	3,0947 ***	0,3581	22,081			
country NL	-0,2287	0,4267	0,796			
country AT	-0,3297	0,3566	0,719			
country PL	2,5769 ***	0,3475	13,157			
country PT	1,4376 **	0,4321	4,210			
country SI	0,1796	0,4516	1,197			
country SK	3,2642 ***	0,4415	26,160			
country FI	-0,5984	0,3536	0,550			
country SE	-0,3938	0,4318	0,675			
country UK	ref.	ref.	ref.			
country IS	-0,4673	0,6692	0,627			
country NO	-0,4939	0,4646	0,610			
Regional-level error terms variances			,			
intercept	0,3818	0,08275				
upper education level in the household	0,2077	0,04486				
number of employed people in the household	0,2198	0,03756				
Regional-level error terms covariances	-,	.,				
COV (intercept, upper education level in the household)	-0,04609	0,04281				
COV (intercept, number of employed people in the HH)	-0,2342	0,05093				
COV (upper education level, nb. of employed people in the HH)	-0,0437	0,03109				
Other parameters	0,0107	0,05107				
Rho coefficient of AR(1)	0,3138	0,002817				
Residual	0,9936	0,002727				
Fit measure: -2 Log Pseudo Likelihood	1832580	0,002727				
Source: EU-SILC data, longitudinal file <sup>30</sup> , 1.08,2009 UDB release.						

Source: EU-SILC data, longitudinal file <sup>30</sup>, 1.08.2009 UDB release, authors' computations. Level of significance for independent variable coefficients: \*: p-value < 0.05; \*\*: p-value < 0.01; \*\*\*: p-value < 0.001

<sup>28</sup> We have used the SAS GLIMMIX command. Useful SAS code examples can be found in Allison (2008) and were kindly given by David Brady.
<sup>29</sup> See appendix 5 for a description of the variables.
<sup>30</sup> 22 countries from this release plus 4 countries from the March release (see above).

Concerning the error terms, SAS/PROC GLIMMIX does not offer a statistical test indicating the level of significance of the variances and covariances of the error terms. But compared to their standard errors, the estimated variances are quite high, which suggests their high level of significance. This in turn justifies on the one hand the choice of the multilevel model, and on the other hand our choice of allowing these variables to have random rather than fixed coefficients. Looking at the empty model (see Appendix 6), we can see that the intra-class correlation (calculated according to the second formula given by Snijders and Bosker, 1999, page 224) is equal to 0.55, meaning that the between-variance is substantial.

Let us examine now the effects of our variables of interest. Recall that our objective is to measure the specific effect of the regional GDP per inhabitant and the regional unemployment rate on the probability of being at-risk-of-poverty. This effect could be either direct or indirect since these macro factors can act through other individual variables on the probability of being poor (such as the education level and the number of employed people in the household).

As expected, the regional GDP per capita has a strong (and highly significant) direct negative effect on the risk of poverty: for individuals living in households where nobody has an upper education level, the odds of being poor (probability of being poor divided by probability of not being poor) decreases by 4.7% (1-0.953=0.047) for an increase of annual GDP per capita by a 1000 Euros. This direct effect is supplemented by an indirect effect: the regional GDP per capita moderates the negative impact of upper education on the poverty risk. In fact, in the average region in terms of GDP per capita (about 24260 Euros/year), the presence of an adult with upper education level decreases the poverty odds by 61% (odds ratio =  $0.39^{31}$ ). In a rich region such as Luxembourg (GDP per capita = 60150 Euros/year), it decreases the odds by only 6.2%; in a quite disadvantaged region like Estonia (GDP per capita = 14547 Euros/year) it decreases the odds by 69 %.

For individuals living in households without any employed people, an additional percentage point of the regional unemployment rate increases by 2.1% the poverty odds, but this effect is only slightly statistically significant (p-value=6%) ceteris paribus (especially when GDP per capita is controlled for). But there is an indirect effect of the regional unemployment rate on the poverty risk, even if rather small: in the average region in terms of the unemployment rate (unemployment rate = 8.52 %), the presence of an additional employed individual decreases the poverty odds by 69%, when controlling for the number of adults in the household (odds ratio =  $0.31^{32}$ ). When the unemployment rate is much lower, such as in Ireland (unemployment rate 4.35%), it decreases these odds by 71%, and by 66% when the unemployment rate is quite high (for example in Slovenia - unemployment rate 14.85%).

To summarize, the moderating effect of the regional unemployment rate on the impact on poverty risk of the number of employed people does exist but it is marginal. By contrast, the moderating effect of the regional GDP per capita on the impact on poverty risk of the presence of highly educated people is quite large.

<sup>&</sup>lt;sup>31</sup> Odds ratio =  $0.39 = \exp(-1.5241 + 0.02427 \times 24.260)$ 

<sup>&</sup>lt;sup>32</sup> Odds ratio =  $0.31 = \exp(-1.2964 + 0.01392 \times 8.52)$ 

Besides and not surprisingly (given the sample size), almost all control variables have an effect on the poverty probability, except for gender<sup>33</sup>: women do not have a higher risk of being poor than men. This can easily be explained by the fact that the poverty status is a household characteristic, which can hardly be influenced by a strictly individual feature<sup>34</sup>. One interesting question is the extent to which these control variables have an impact on the poverty risk (even if our study focuses on the possible impact of macro determinants on the effect some factors of interest can have on the poverty risk):

- if the activity of at least one household member is hampered by disease, the odds of being poor increase by 18%

- controlling for especially the fact that the activity of at least one household member is hampered (or not) by disease, the chronic character of this disease decreases by 11% the odds of being poor (this could be a consequence of the social benefits people with chronic diseases are entitled to)

- the number of children in the household has a slightly concave effect on the odds of being poor: as the number of children increases, the effect of an additional child decreases progressively up to a value (14 children) lying beyond the observed maximum in the sample (12 children). But the effect of each additional child remains substantial. As an example, while a first child increases by 41% the odds of being poor, a fourth child still increases them by 31%

- the number of adults in the household also proves to have a clear non linear effect on the poverty risk: ceteris paribus (and especially when controlling for the number of employed people), the odds of being poor are virtually the same if there are one or two adults in the household, but they increase by 21% with the third adult and by 45% with the fourth one.

#### 6. Conclusion

Analysing the determinants of the monetary poverty probability has already been attempted by many studies. But few of them have simultaneously used panel data, considered factors at the macro level, and used the right techniques to deal with all these elements.

As for our results, they show that both the regional GDP per capita and the regional unemployment rate do have an effect on poverty risk.

<sup>&</sup>lt;sup>33</sup> Since the poverty status is defined at the household level, it could be considered as irrelevant to introduce variables measured at individual level such as age and gender. We have added these two variables to our model for comparison's sake, as almost all studies do this as well. As far as age is concerned, it could be argued that, even though age is, like gender, a factor measured at the individual level, it might have some signification as an household characteristic: due to frequent endogamy, the age of one adult belonging to the household offers some clue about the age of other adults in the household, at least for non single adult households. With these limitations in mind, we can interpret the (highly significant) quadratic effect of age: ceteris paribus, the poverty risk is first decreasing more than linearly with age, reaching a minimum at age 44, and then increasing more than linearly. For example, at age 25, an additional year results in a decrease of the poverty risk by 1.2%; and at age 50, an additional year of age results in an increase of 1.1%.

<sup>&</sup>lt;sup>34</sup> This is the very reason why the different common indicators agreed upon in the context of the OMC on Social Protection and Social Inclusion which focus on gender differences are calculated on single person households.

In terms of economic and social policy implications, it means that:

- policies oriented towards higher economic growth rates in disadvantaged European regions are able to alleviate the risk of poverty even if poverty is defined in relative terms

- this kind of economic policies, if successful in its effort to sustain the economic well-being of families in poor regions, will, as a side-effect, diminish the anti-poverty effect of the presence in the household of higher educated people. We suspect that this indirect effect is associated to the choice of defining poverty as a relative concept – a European view, which countries like the US do not share

- as for the regional unemployment rate, its direct positive impact on the poverty risk is essentially a confirmation of what could be expected and of what is already known, even if the weakness of this effect is quite surprising. Still more surprising is the fact that the regional unemployment rate does not moderate to a large extent the impact on the poverty risk of the number of employed people in the household. This implies that policies aimed at combating high regional unemployment rates will, as such, unfortunately not lower to a large extent the regional poverty rates.

However, our analysis faces two types of limitations. The first one results from methodological choices we have made, the second is due to the data.

First, because we needed an indicator differentiating and thus ranking the 93 regions in terms of poverty rates, we have made use of a European poverty threshold, which has proven quite relevant in terms of ability to estimate the econometric model. However, precisely because the European regions are quite dispersed around the average at-risk-of-poverty rate, we were not able to check the consistency of the results in using alternative measures of the European threshold<sup>35</sup> (such as 50% or 70% of the European median equivalent income).

Concerning then the data, and with the results at country-level in mind, we would have liked to test the effect of other macro characteristics at the regional level, such as the expenses in unemployment or social benefits (expressed in percentage of the GDP). But these were not available at the regional level. In future analyses, we would thus be interested in adding some variables of that kind, once they are available. And, as for the regions, the variable defined in the EU-SILC dataset is missing in some countries (even in quite big countries such as Germany and the United Kingdom). In order to keep these countries in the analysis, we have defined each of them as a unique – and quite large – region.

<sup>&</sup>lt;sup>35</sup> In fact, with a poverty threshold equal to 60% of the European median equivalent income, the at-risk-ofpovery rates of the different countries range between 1% and 82% in 2005, and between 2% and 75% in 2006. Changing this threshold for a lower (higher) one would lead to even lower (higher) rates in the richest (poorest) countries, making the analysis impossible to run.

#### References

ALLISON Paul D. (2008), Longitudinal Data Analysis Using SAS, October, 100 pages.

**ANDRIOPOULOU Eirini**, **TSAKLOGLOU Panos** (2008), "Once poor, always poor? Do initial conditions matter? Evidence from the ECHP", IRISS Conference, Oct., 33 pages.

**ANGRIST Joshua D., PISCHKE Jörn-Steffen** (2009), "Mostly Harmless Econometrics. An Empiricist's Companion", Princeton University Press, 373 p.

**AYLLON Sara** (2008), "Modelling poverty transitions in Spain: Do attrition and initial conditions really matter?", IRISS Working Paper Series, 2008-08, Oct., 36 pages.

**BRADBURY Bruce**, JÄNTTI Markus (1999) "Child Poverty across Industrialised Nations", Innocenti Occasional Papers, Economic and Social Policy Series, UNICEF, n°71.

**BRADSHAW Jonathan** (1999), "Child Poverty in Comparative Perspective", *European Journal of Social Security*, Vol.1, pp. 383-406.

**BRADY David**, **FULLERTON Andrew S.**, **MOREN-CROSS Jennifer** (2009), "Putting Poverty in Political Context: A Multi-Level Analysis of Adult Poverty Across 18 Affluent Countries", *Social Forces*, Vol. 88(1), September, pp. 271-300.

**BUDDELMEYER Hielke**, **VERICK Sher** (2007), "Understanding the Drivers of Poverty Dynamics in Australian Households", IZA DP No. 2827, 29 pages.

CAPPELLARI Lorenzo (2004), "Earnings Mobility Among Italian Low Paid Workers", *Journal of Population Economics*, Vol. 20(3), pp. 465-482.

**CAPPELLARI Lorenzo**, **JENKINS Stephen P**. (2002a), "Who Stays Poor? Who Becomes Poor? Evidence from the British Household Panel Survey", *The Economic Journal*, Vol. 112, No. 478, Conference Papers (Mar., 2002), pp. C60-C67.

**CAPPELLARI Lorenzo**, **JENKINS Stephen P.** (2002b), "Modelling Low Income Transitions", *ISER Working Papers Number 2002-8*, 42 pages.

**CAPPELLARI Lorenzo**, **JENKINS Stephen P.** (2004), "Modelling Low Income Transitions", *Journal of Applied Econometrics, Vol.* 19, pp. 593–610.

**FERTIG Michael, TAMM Marcus** (2007), "Always Poor or Never Poor and Nothing in Between? Duration of Child Poverty in Germany", IZA Discussion Paper N° 2645, February, 29 pages.

**MARLIER Eric**, **ATKINSON A.B.**, **CANTILLON Bea**, **NOLAN Brian** (2007), The EU and Social Inclusion – Facing the challenges, The Policy Press, The University of Bristol, 303 pages.

**MEJER Lene**, **SIERMANN Clemens** (2000), "La pauvreté monétaire dans l'Union européenne : la situation des enfants, les différences entre les sexes et l'écart de pauvreté", Eurostat, *Statistiques en bref*, Population et Conditions sociales, 12/2000, 7 pages.

**MOLLER Stephanie, HUBER Evelyne, STEPHENS John D., BRADLEY David, NIELSEN François** (2003), "Determinants of Relative Poverty in Advanced Capitalist Democracies", *American Sociological Review*, Vol. 68, No. 1, February, pp. 22-51.

**MOON Choon-Geol, STOTSKY Janet G.** (1993), "The Effect of Rent Control on Housing Quality Change: A Longitudinal Analysis", *The Journal of Political Economy*, Vol. 101, No. 6, December, pp. 1114-1148.

**OECD** (2001), "When Money is Tight: Poverty Dynamics in OECD Countries", in *Employment Outlook*, Chapter 2, pp. 37-87.

**POGGI Ambra** (2007), "Does persistence of social exclusion exist in Spain?", *Journal of Economic Inequality*, Vol. 5, pp. 53-72.

**RAY Jean-Claude**, **JEANDIDIER Bruno** (2003), "Depuis une décennie, que nous ont appris les données longitudinales à propos de la pauvreté en France? Une première synthèse", Communication aux X<sup>es</sup> Journées du Longitudinal, Caen, mai, 61 pages.

**SNIJDERS Tom A.B., BOSKER Roel J.** (2004), *Multilevel Analysis. An introduction to basic and advanced multilevel modeling*, Sage Publications, 266 pages.

STEWART Mark B., SWAFFIELD Joanna K. (1999), "Low Pay Dynamics and Transition Probabilities", *Economica*, Vol. 66, pp. 23-42.

**TAI Tsui-o**, **TREAS Judith** (2008), "Poverty, Household Composition, and Welfare States: A Multi-level Analysis of 22 Countries", Luxembourg Income Study Working Paper Series, Working Paper No. 492, September, 33 pages.

**UNICEF** (2000), "A League Table of Child Poverty in Rich Nations", Innocenti Report Card, Issue n°1, June, 28 pages.

**VAN KERM Philippe** (2004), "Une évaluation économétrique des flux vers et hors de la pauvreté en Belgique", IRISS Working Paper Series, 2004-04, 17 pages.

**WIEPKING Pamala**, **MAAS Ineke** (2005), "Gender Differences in Poverty: A Cross-National Study", *European Sociological Review*, Vol. 21(3), July, pp. 187-200.

#### Appendix 1. At-risk-of-poverty rates of working aged adults (25-55) in Europe

income)		
Country	2005	2006
AT	5	6
BE	6	8
CY	7	6
CZ	42	41
DE	7	6
DK	3	3
EE	70	64
ES	20	19
FI	5	6
FR	8	9
EL	24	26
HU	73	69
IE	8	8
IS	3	3
IT	14	17
LT	79	75
LU	1	2
LV	81	74
NL	6	3 3
NO	3	
PL	74	73
PT	42	40
SE	5	5
SI	13	13
SK	77	72
UK	7	8

At-risk-of-poverty rates of working age adults (25-55) in Europe (European threshold = 60% of the European median equivalent income)

Source: EU-SILC data, longitudinal file, 1.08.2009 UDB release, authors' computations.

Reading note: with the European poverty threshold calculated for the whole population <sup>36</sup>, 5% of individuals aged 25-55 in Austria were at-risk-of-poverty in 2005.

We can notice that the countries face very different situations in terms of the at-riskof-poverty rate, a conclusion that cannot be drawn from the figures shown in the table below (based on national poverty thresholds).

<sup>&</sup>lt;sup>36</sup> The European poverty threshold is calculated by taking into account all individuals living in the 26 countries under study. In other words, children and elderly people are not excluded from this calculation, even if they are not kept in the analyses afterwards.

Country	2005	2006
AT	11	11
BE	11	11
CY	10	10
CZ	10	9
DE	11	12
DK	9	9
EE	16	15
ES	16	15
FI	8	9
FR	10	11
EL	16	17
HU	14	15
IE	14	14
IS	9	8
IT	16	17
LT	19	18
LU	13	14
LV	18	19
NL	10	8
NO	8	8
PL	21	19
PT	15	15
SE	7	10
SI	10	9
SK	13	11
UK	14	14

At-risk-of-poverty rates of working aged adults (25-55) in Europe (national thresholds)

Source: EU-SILC data, cross-sectional file, 1.08.2009 UDB release. Reading note: with the poverty threshold calculated at national level, 11% of individuals aged 25-55 in Austria were at-risk-of-poverty in 2005.

These official figures are based on the cross-sectional file, which means that all individuals are taken into account (whereas the longitudinal file concerns only those present at least two years).

## Appendix 2. Poverty probability determinants at country-level

		Deeper	an Drafila			
	Ord	dered	se Profile	Total		
	7		ndicator 1			
		1 1 2 0		72218 211983		
The G	LIMMIX procedure :		e probabilit <sup>,</sup>		dicator='1'	·.
	*	-				
	C	Dim side Cov. Par	ensions	3		
		-side Cov. Par -side Cov. Par		2		
	Co	olumns in X		21		
		olumns in Z pe		3		
		ubjects (Block ax Obs per Sub		26 30703		
		-	-			
	Optimi	Optimizati zation Techniq	on Informatio	on vton-Raphson		
		ters in Optimi		veon naphson		
	Lower H	Boundaries	4			
		Boundaries Effects	1	ofiled		
		al Variance		ofiled		
		ng From	GLI	4 estimates		
/ 32 iterations		e criterion (P	CONV=1 11022	(-8) estictio	d	
	convergence		tatistics	5 0) Sacistie	u.	
		es Log Pseudo-				
		ralized Chi-Sq r. Chi-Square		291051.3 1.02		
	Gene	. chi bquare	/ DE	1.02		
		Covariance Pa	rameter Estir		+	
	Cov Parm	Subject	T	S Stimate	tandard Error	
	Intercept	country			0.1091	
	uppereducHH	country		0.2754 0.1406	0.08766	
	nbemployedHH AR(1)			0.1406		
	Residual			1.0242 0		
	Normatotic Corre-	niango Matrix	of Coursiana	Domomotor F	atimataa	
Cov Parm	Asymptotic Cova: Subject	Cov	Pl Covi	2 CovP	3 Cot	7P4 CovP5
Intercept	country	0.011	90 0.00008	39 -0.0001	9 -6.96E	E-7 -5.7E-7
uppereducHH	country	0.0000	89 0.00768	34 0.00003	2 -6.59E	E-7 -3.03E-7
nbemployedHH AR(1)	country country ID_unique_UE27(co	-0.000 ountry) -6.96E	-7 -6.59E	-7 3.937E-	7 7.707E	E-7 2.309E-8 E-6 2.001E-6
Residual		-5.7E	-7 -3.03E	-7 2.309E-	8 2.001E	E-6 7.903E-6
	Asymptotic Corre	lation Matrix	of Covariance	Parameter E	stimates	
Cov Parm	Subject	Cov	Pl Covi	2 CovP	3 Cot	
Intercept	country	1.00	00 0.00933	33 -0.0377	6 -0.002	230 -0.00186 271 -0.00123
uppereducHH nbemployedHH	country	-0.037	76 0.0082	29 1.000	9 -0.002	271 -0.00123 152 0.000183
AR(1)	country ID_unique_UE27(co	ountry) -0.002				
Residual		-0.001	86 -0.0012	23 0.00018	3 0.25	564 1.0000
		Solutions fo	r Fixed Effe	cts		
			Standard			
Effect Intercept	wave	e Estimate 1.4742			t Value 3.11	Pr >  t  0.0046
wave	200			284E3	-6.81	<.0001
wave	200			•	•	•
GDPhabnuts GDPvariati		-0.1103 0.1560		284E3 284E3	-7.68 7.97	<.0001 <.0001
unempraten		0.1386		284E3	9.56	<.0001
socexclexp		-0.7600		284E3	-5.88	<.0001
unemployme woman	entexpenses	-0.2225 0.01188		284E3 284E3	-2.14 0.85	0.0327 0.3949
age_center	ed	0.001467		284E3	1.72	0.0849
age_center		0.000469		284E3	4.46	<.0001
chronicdis activityha		-0.1419 0.1720		284E3 284E3	-8.38 9.81	<.0001 <.0001
uppereducH	*	-1.2942		25415	-9.54	<.0001
	enses*uppereducHH	0.7339		284E3	4.02	<.0001
nbchildren nbchildren		0.3795 -0.01734		284E3 284E3	24.82 -4.06	<.0001 <.0001
nbadultsHH		-0.2065		284E3	-7.46	<.0001
nbadultsHH	12	0.08637	0.004631	284E3	18.65	<.0001
nbemployed	lHH entexpenses*nbemplo	-1.0282 ovedHH -0 1153		25 284E3	-9.08 -1.74	<.0001 0.0812
απειιρτογίιε	meerbenses incellbr(	-v.1133	0.00013	20100	1./4	0.0012

## Appendix 3. Sample size, by country and year

Country	2005	2006
AT	3882	5469
BE	2408	4002
CY	2357	3373
CZ	4391	7367
DE	10529	9285
DK	3549	3451
EE	2665	4269
ES	8472	11501
FI	4021	5528
FR	5876	7324
EL	4642	4211
HU	4353	6609
IE	2872	1992
IS	1584	2243
IT	12955	18048
LT	2533	3713
LU	4322	4601
LV	2704	3438
NL	8574	9287
NO	4088	3926
PL	10627	14644
PT	2671	3662
SE	3541	4751
SI	8671	11271
SK	3447	4898
UK	6224	7594

Number of individuals aged 25-55, in each country, for each year (sample size, unweighted cases)

Source: EU-SILC data, longitudinal file, 1.08.2009 UDB release, authors' computations.

# Appendix 4. List of available regions<sup>37</sup> in the dataset

Austria		ES62	Región de Murcia
AT1	Ostösterreich	ES63	Ciudad Autónoma de Ceuta
AT2	Südösterreich	ES64	Ciudad Autónoma de Melilla
AT3	Westösterreich	ES70	Canarias
Belgium		Finland	
BE1	Région de Bruxelles-Capitale /	FI13	Itä-Suomi
	Brussels Hoofdstedelijk Gewest	FI18	Etelä-Suomi
BE2	Vlaams Gewest	FI19	Länsi-Suomi
BE3	Région Wallonne	FI1A	Pohjois-Suomi
Republic of Cy	nrus	France	
CY0	Kypros / Kibris	FR10	Île de France
010	Kypios / Kions	FR21	Champagne-Ardenne
Czech Republi	e de la companya de l	FR22	Picardie
CZ01	Praha	FR23	Haute-Normandie
CZ01 CZ02		FR24	Centre
CZ02 CZ03	Stredni Cechy	FR24 FR25	Basse-Normandie
	Jihozapad Savarazarad	-	
CZ04	Severozapad	FR26	Bourgogne
CZ05	Severovychod	FR30	Nord - Pas-de-Calais
CZ06	Jihovychod	FR41	Lorraine
CZ07	Stredni Morava	FR42	Alsace
CZ08	Moravskoslezsko	FR43	Franche-Comté
~		FR51	Pays de la Loire
Germany		FR52	Bretagne
			Poitou-Charentes
longitudinal dataset for Germany.		FR61	Aquitaine
DE	Germany	FR62	Midi-Pyrénées
		FR63	Limousin
Denmark		FR71	Rhône-Alpes
DK0	Denmark	FR72	Auvergne
		FR81	Languedoc-Roussillon
Estonia		FR82	Provence-Alpes-Côte d'Azur
EE0	Estonia	FR83	Corse
Spain		Greece	
ES11	Galicia	GR1	Voreia Ellada
ES12	Principado de Asturias	GR2	Kentriki Ellada
ES13	Cantabria	GR3	Attiki
ES21	País Vasco	GR4	Nisia Aigaiou, Kriti
ES22	Comunidad Foral de Navarra		C A
ES23	La Rioja	Hungary	
ES24	Aragón	HU1 .	Kozep-Magyarorszag
ES30	Comunidad de Madrid	HU2	Dunantul
ES41	Castilla y León	HU3	Alfold Es Eszak
ES42	Castilla-La Mancha		
ES43	Extremadura	Ireland	
ES51	Cataluña		able is not available in the
ES52	Comunidad Valenciana	e e	taset for Ireland.
ES52 ES53	Illes Balears	IE0	Ireland
ES61	Andalucía		nomulu
L001	/ manucia	Iceland	
		IS	Iceland

<sup>&</sup>lt;sup>37</sup> Details from EU-SILC documentation and http://ec.europa.eu/eurostat/ramon/nuts/codelist\_en.cf m?list=nuts

Italy		Poland	
ITC	Nord-Ovest	PL1	Region Centralny
ITD	Nord-Est	PL2	Region Poludniowy
ITE	Centro (I)	PL3	Region Wschodni
ITF	Sud	PL4	Region Polnocno-Zachodni
ITG	Isole	PL5	Region Poludniowo-Zachodni
		PL6	Region Polnocny
Lithuania			
LT0	Lietuva	Portugal	
		PT	Portugal
Luxembourg			-
LU0	Luxembourg (Grand-Duché)	Sweden	
		SE	Sweden
Latvia			
LV0	Latvija	Slovenia	
		SI	Slovenia
The Netherland	l		
The region varia	ble is not available in the	Slovakia	
	aset for the Netherlands.	SK0	Slovenska
NL	The Netherlands		
		The United Kin	ngdom
Norway			able is not available in the
NO0	Norway	longitudinal dat	
	-	UK	United-Kingdom

## Appendix 5. Description of the explanatory variables and descriptive statistics

anatory variables	
Label of the variable	Description of the variable
woman	EU-SILC variable RB090; woman=1 if RB090=2
age centered (around the	EU-SILC variable RX020, centered (age in the year
average: 41.27)	prior to the survey)
chronic disease in the	Authors' calculations using the EU-SILC variable
household	PH020: is there at least one household member who
	suffers from a chronic disease?
activity hampered by disease in	Authors' calculations using the EU-SILC variable
the household	PH030: is there at least one household member
	whose activities are hampered because of health
	problems?
upper education level in the	Authors' calculations using the EU-SILC variable
household	PE040: is there at least one household member
	whose upper level of education is tertiary education
	(PE040=5)?
number of children in the	number of children (age 0-14) in the year prior to the
household	survey
number of adults in the	number of adults (age 18 or more) in the year prior to
household	the survey
number of employed people in	number of employed household members (authors'
the household	calculations using the EU-SILC variable PL030 -
	codes 1 or 2)
country	EU-SILC variable RB020
wave	EU-SILC variable RB010
regional annual GDP per capita	Information from Eurostat
(in 10 <sup>3</sup> Euros)	
regional unemployment rate	Information from Eurostat
(expressed in %)	
	Label of the variable         woman         age centered (around the average: 41.27)         chronic disease in the household         activity hampered by disease in the household         upper education level in the household         number of children in the household         number of adults in the household         number of employed people in the household         country         wave         regional annual GDP per capita (in 10 <sup>3</sup> Euros)         regional unemployment rate

#### Description of the explanatory variables

### Descriptive statistics for the whole sample

Wave	N Obs	Variable	Ν	NMiss	Mean	Std Dev	Minimum	Maximum
2005	131891	pov indicator	131891	0	0.1965	0.4259	0	1.0000
		woman	131887	4	0.5130	0.5357	0	1.0000
		age centered	131891	0	-0.6425	9.1801	-16.2665	14.7335
		age_centered2	131891	0	73.7935	76.2358	0.0710	264.6
		chronicdiseaseHH	131891	0	0.3885	0.5223	0	1.0000
		activityhamperedHH	131891	0	0.2931	0.4878	0	1.0000
		uppereducHH	126795	5096	0.3962	0.5182	0	1.0000
		nbchildren	131891	0	0.9077	1.1426	0	11.0000
		nbadultsHH	131891	0	2.3903	1.0792	1.0000	10.0000
		nbemployedHH	129384	2507	1.5131	0.8971	0	8.0000
		GDPhabnuts	131891	0	23.7538	7.1961	8.2000	57.1000
		unempratenuts	131891	0	8.8728	4.4381	2.5000	21.4000
2006	166379	pov_indicator	166379	0	0.2000	0.3757	0	1.0000
		woman	166371	8	0.5128	0.4695	0	1.0000
		age_centered	166379	0	-0.2041	7.9962	-16.2665	14.7335
		age_centered2	166379	0	72.5178	66.5149	0.0710	264.6
		chronicdiseaseHH	166379	0	0.3911	0.4584	0	1.0000
		activityhamperedHH	166379	0	0.2926	0.4273	0	1.0000
		uppereducHH	158892	7487	0.4059	0.4560	0	1.0000
		nbchildren	166379	0	0.9288	0.9917	0	12.0000
		nbadultsHH	166379	0	2.4239	0.9506	1.0000	11.0000
		nbemployedHH	162220	4159	1.6016	0.7713	0	8.0000
		GDPhabnuts	166379	0	24.7924	6.6164	8.7000	63.1000
		unempratenuts	166379	0	8.1568	3.0558	2.8000	21.0000

## Descriptive statistics for at country level

wave	country	N Ob	s Variable		N Miss		Std Dev	7 Minimum	Maximum
2005	BE	2408	pov_indicator	2408	0	0.0609		0	1.0000
			woman	2408		0.4920		0	1.0000
			age_centered	2408 2408	0	0.2142			14.7335
			age_centered2	2408	0	72.2531		0.0710	264.6
			chronicdiseaseHH	2408 2408	0	0.3267		0	1.0000
			activityhamperedHH	2408 2089	210	0.3100		0	1.0000
			uppereducHH nbchildren	2408	319 0	0.5108		0	1.0000 7.0000
				2408	0	0.9201		1.0000	7.0000
			nbadultsHH nbemployedHH	2408	0	2.3255 1.4716		1.0000	5.0000
			GDPhabnuts	2408	0	26.6319		19.4000	53.3000
			unempratenuts	2408	0	8.4371		5.4000	16.3000
006	BE	4002	pov_indicator	4002	0	0.0789	0.2398	0	1.0000
			woman	4002	0	0.4979		0	1.0000
			age_centered	4002		0.0482		-16.2665	14.7335
			age_centered2	4002		74.3560		0.0710	264.6
			chronicdiseaseHH	4002		0.3407		0	1.0000
			activityhamperedHH	4002	0	0.3010		0	1.0000
			uppereducHH	3634		0.5390		0	1.0000
			nbchildren	4002		0.8937		0	7.0000
			nbadultsHH	4002	0	2.2926		1.0000	7.0000
			nbemployedHH	3969	33	1.5155		0	4.0000
			GDPhabnuts	4002		28.0513		20.1000	55.1000
			unempratenuts	4002	0	8.3091	3.8080	5.0000	17.6000
005	CZ	4391		4391	0	0.4156		0	1.0000
			woman	4391	0	0.4986		-16 2665	1.0000
			age_centered	4391		-1.3598		-16.2665	14.7335
			age_centered2	4391 4391		88.9707		0.0710	264.6 1.0000
			chronicdiseaseHH activityhamperedHH	4391 4391		0.3762 0.3349		0	1.0000
				4391	0			0	1.0000
			uppereducHH nbchildren	4391	0	0.2071 0.7598		0	5.0000
			nbadultsHH	4391	0	2.5584			7.0000
			nbemployedHH	4391	0	1.7005		00001	4.0000
			GDPhabnuts	4391	0	17.0978		13.3000	35.6000
			unempratenuts	4391	0	7.9891	3.1642	3.5000	13.9000
006	CZ	7367	pov indicator	7367	0	0.4136	0.3314	0	1.0000
			woman	7367	0	0.5002		0	1.0000
			age centered	7367	0	-1.2929		-16.2665	14.7335
			age centered2	7367	0	87.7967		0.0710	264.6
			chronicdiseaseHH	7367	0	0.3807	0.3267	0	1.0000
			activityhamperedHH	7367	0	0.3379	0.3183	0	1.0000
			uppereducHH	7323	44	0.2118	0.2749	0	1.0000
			nbchildren	7367	0	0.7563	0.6111	0	7.0000
			nbadultsHH	7367	0	2.5852		1.0000	8.0000
			nbemployedHH	7367	0	1.7118	0.6076	0	5.0000
			GDPhabnuts	7367	0	18.2129	4.8756	14.2000	38.4000
			unempratenuts	7367	0	7.2759	2.1418	2.8000	12.8000
005	DK	3549	pov_indicator	3549	0	0.0331	0.1237	0	1.0000
			woman	3549		0.4989	0.3458		1.0000
			age_centered	3549	0	-0.1867		-16.2665	14.7335
			age_centered2	3549	0	74.1813	48.3555	0.0710	264.6
			chronicdiseaseHH	3549	0	0.2248	0.2887	0	1.0000
			activityhamperedHH	3549 3488	0	0.1249	0.2287 0.3413	0	1.0000
			uppereducHH	3488	61	0.4311		0	1.0000
			nbchildren	3549	0	0.9719	0.7503 0.4682	0	7.0000 7.0000
			nbadultsHH	3549 3491	0 58	1.9228	0.4682 0.4816	1.0000	
			nbemployedHH GDPhabnuts	3491 3549	58 0	1.5338 27.8000	0.4816	0 27.8000	4.0000 27.8000
			unempratenuts	3549	0	4.8000	0	4.8000	4.8000
006	DK	3451	pov indicator	3451	0	0.0320	0.1208	0	1.0000
			woman	3451	0	0.4933	0.3430	0	1.0000
			age centered	3451	0	-0.0156		-16.2665	14.7335
			age centered2	3451	0	69.9331	46.3429	0.0710	264.6
			chronicdiseaseHH	3451	0	0.2375	0.2919	0	1.0000
			activityhamperedHH	3451	0	0.1280	0.2292	0	1.0000
			uppereducHH	3391	60	0.4392	0.3390	0	1.0000
			nbchildren	3451	0	0.9772	0.7343	0	5.0000
			nbadultsHH	3451	0	1.9430	0.4724	1.0000	5.0000
			nbemployedHH	3382	69	1.5658	0.4818	0	4.0000
			GDPhabnuts	3451	0	29.1000	0	29.1000	29.1000
			unempratenuts	3451	0	3.9000	0	3.9000	3.9000
	DE	10529	pov_indicator	10529	0	0.0748	0.4145	0	1.0000
005	55								
05	01		woman	10529	0	0.5332	0.7863	0	1.0000
05			age_centered	10529	0	-0.2368	13.1238	-16.2665	14.7335
005									

			activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	10529 10364 10529 10529 10529 10529 10529	0 165 0 0 0 0	0.4231 0.5070 0.7408 2.1375 1.3698 26.3000 11.1000	0.7786 0.7878 1.5855 1.3513 1.2661 0 0	0 0 1.0000 0 26.3000 11.1000	1.0000 1.0000 6.0000 7.0000 5.0000 26.3000 11.1000
2006	DE	9285	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	9285 9285 9285 9285 9285 9285 9285 9285	0 0 0 0 0 0 0 0 0 0 0 0 0	0.0624 0.5655 1.8870 63.1927 0.4929 0.3678 0.5429 0.9117 2.2395 1.6796 27.4000 10.2000	0.3895 0.7983 12.4357 108.7 0.8051 0.7766 0.8022 1.6225 1.4165 1.2261 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 27.4000 10.2000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 1.0000\\ 6.0000\\ 6.0000\\ 5.0000\\ 27.4000\\ 10.2000\end{array}$
2005	EE	2648	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	2648 2648 2648 2648 2648 2648 2648 2648		0.6987 0.5190 -0.8065 81.8660 0.4745 0.4575 0.4219 0.8302 2.4462 1.6159 13.7000 7.9000	0.1792 0.1952 3.5205 28.6931 0.1951 0.1946 0.1929 0.3793 0.4073 0.3584 0 0	0 0 -16.2665 0.0710 0 0 1.0000 0 1.0000 0 13.7000 7.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 8.0000 6.0000 13.7000 7.9000
2006	EE	4257	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	4257 4257 4257 4257 4257 4257 4257 4257	0 0 0 0 30 0 0 0 0 0 0	0.6417 0.5192 -0.8308 81.1909 0.4938 0.4280 0.4447 0.7761 2.4614 1.6337 15.4000 5.9000	0.1471 0.1533 2.7528 22.5456 0.1534 0.1518 0.1525 0.2906 0.3270 0.2754 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 1.0000 0 15.4000 5.9000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 9.0000\\ 8.0000\\ 6.0000\\ 15.4000\\ 5.9000\\ \end{array}$
2005	IE	2872	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	2872 2872 2872 2872 2872 2872 2872 2872	0 0 0 95 0 0 0 0 0	0.0824 0.5097 -0.2720 75.3925 0.3453 0.2946 0.4414 1.1631 2.5971 1.6651 32.4000 4.3000	0.1751 0.3183 5.5261 44.0678 0.3027 0.2903 0.3168 0.7583 0.7238 0.6233 0 0 0	0 0 -16.2665 0.0710 0 0 1.0000 0 32.4000 4.3000	1.0000 1.0000 13.7335 264.6 1.0000 1.0000 9.0000 9.0000 6.0000 32.4000 4.3000
2006	IE	1992	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	1992 1992 1992 1992 1992 1992 1992 1992	0 0 0 63 0 0 0 0 0	0.0811 0.5266 0.5228 72.0837 0.3667 0.2901 0.4180 1.2204 2.5523 1.5983 34.8000 4.4000	0.2023 0.3700 6.2803 50.2126 0.3571 0.3363 0.3660 0.9358 0.7821 0.6942 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 34.8000 4.4000	1.0000 1.0000 13.7335 264.6 1.0000 1.0000 9.0000 7.0000 7.0000 5.0000 34.8000 4.4000
2005	EL	4642	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	$\begin{array}{c} 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4446\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ 4642\\ \end{array}$	0 0 0 0 196 0 0 0 0 0 0	0.2368 0.5011 -1.1845 77.5079 0.2578 0.2381 0.3357 0.7773 2.6662 1.5917 20.8883 9.9011	0.3385 0.3981 6.9457 57.6718 0.3483 0.3391 0.3777 0.7372 0.7512 0.6483 4.4502 0.9685	0 0 -16.2665 0.0710 0 0 0 1.0000 0 1.0000 0 16.2000 8.2000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 1.0000 7.0000 28.3000 11.4000

2006	EL	4211	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	4211 4211 4211 4211 4211 4027 4211 4211 4211 4211 4211	0 0 0 0 184 0 0 0 0 0	0.2574 0.5024 -0.8390 76.5434 0.2626 0.2301 0.3405 0.7931 2.6953 1.6258 21.8582 8.9479	0.3462 0.3960 6.8965 57.1113 0.3485 0.3333 0.3772 0.7434 0.7519 0.6779 4.8696 0.5389	0 0 -16.2665 0.0710 0 0 1.0000 0 1.0000 0 16.9000 7.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 0.0000 7.0000 30.5000 9.7000
2005	ES	8429	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	8429 8429 8429 8429 8429 8429 8429 8429	0 0 0 635 0 0 0 0 0	0.2004 0.4927 -1.7718 78.3910 0.3623 0.3263 0.4688 0.7216 2.6879 1.6190 23.1034 9.2296	0.5233 0.6537 11.3422 95.7980 0.6285 0.6131 0.6525 1.2211 1.4681 1.1830 6.0902 3.6424	0 0 -16.2665 0.0710 0 0 0 1.0000 0 15.6000 5.6000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 8.0000 8.0000 29.9000 19.7000
2006	ES	11438	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	11438 11438 11438 11438 11438 11438 10605 11438 11438 11438 11438 11438	0 0 0 833 0 0 0 0 0	0.1913 0.4098 -1.8626 78.3685 0.3500 0.3440 0.4680 0.7259 2.6490 1.6282 24.6383 8.6267	0.4309 0.5387 9.4802 80.1413 0.5225 0.5204 0.5469 0.9986 1.1568 0.9546 5.4181 2.7513	0 0 -16.2665 0.0710 0 0 0 1.0000 0 16.7000 5.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 11.0000 7.0000 32.1000 21.0000
2005	FR	5876	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	5876 5876 5876 5876 5876 5876 5876 5876	0 0 0 0 177 0 0 2 0 0	0.0798 0.5252 1.3631 67.8346 0.4439 0.2686 0.3917 1.0526 2.2903 1.5306 25.2810 8.8630	0.4672 0.8610 14.0051 116.5 0.7642 0.8420 1.9390 1.5338 1.2917 11.6435 3.0632	0 0 -16.2665 0.0710 0 0 0 1.0000 0 19.6000 6.4000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 7.0000 7.0000 4.0000 38.7000 13.2000
2006	FR	7324	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	7324 7324 7324 7324 7324 7324 7324 7324	0 0 0 205 0 0 0 0 0 0	$\begin{array}{c} 0.0917\\ 0.5262\\ 1.2819\\ 68.6154\\ 0.4242\\ 0.2657\\ 0.3899\\ 1.0192\\ 2.2675\\ 1.5391\\ 26.1066\\ 8.8601 \end{array}$	0.4396 0.7608 12.4687 104.2 0.7530 0.6730 0.7433 1.6985 1.3719 1.1152 10.5634 2.5799	0 0 -16.2665 0.0710 0 0 0 1.0000 20.3000 6.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 7.0000 7.0000 4.0000 4.0000 12.4000
2005	IT	12955	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	12955 12955 12955 12955 12955 12955 12955 12955 12955 12955 12955 12955	0 0 0 0 158 0 0 0 0 0	0.1430 0.4996 -0.6673 73.6183 0.2886 0.2261 0.2134 0.7360 2.5511 1.4667 23.6233 8.1620	0.4209 0.6012 10.2847 83.8161 0.5448 0.5030 0.4925 1.0747 1.2129 0.9832 7.0590 5.5057	0 0 -16.2665 0.0710 0 0 1.0000 0 1.0000 0 15.6000 4.0000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 5.0000 5.0000 28.8000 15.3000
2006	IT	18048	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH</pre>	18048 18048 18048 18048 18048 18048 17906 18048 18048 18048	0 0 0 0 0 142 0 0 0	0.1718 0.4995 -0.4297 72.2801 0.2796 0.2694 0.2191 0.7350 2.5369 1.4865	0.3850 0.5103 8.6658 70.6271 0.4581 0.4528 0.4222 0.9174 1.0417 0.8410	0 0 -16.2665 0.0710 0 0 0 1.0000 0	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 1.0000\\ 6.0000\\ 7.0000\\ 6.0000\\ \end{array}$

			GDPhabnuts unempratenuts	18048 18048	0 0	24.5302 7.1662	6.1430 3.8868	16.3000 3.6000	29.8000 12.7000
2005	СҮ	2357	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	2357 2357 2357 2357 2357 2357 2357 2357	0 0 0 0 42 0 0 0 0 0 0	0.0719 0.5104 -1.1817 80.6322 0.3778 0.3540 0.4840 0.9912 2.7608 1.8203 20.4000 5.3000	0.0821 0.1588 2.8284 23.5761 0.1520 0.1589 0.3387 0.3348 0.2727 0 0 0	0 0 -16.2665 0.0710 0 0 1.0000 0 20.4000 5.3000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 1.0000\\ 8.0000\\ 7.0000\\ 6.0000\\ 20.4000\\ 5.3000\end{array}$
2006	СҮ	3373	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	3373 3373 3373 3373 3373 3373 3319 3373 3373	0 0 0 0 54 0 0 0 0 0	0.0552 0.5142 -1.0617 80.5458 0.4170 0.2706 0.4900 1.0275 2.8076 1.8817 21.3000 4.5000	0.0601 0.1316 2.3468 19.5748 0.1298 0.1170 0.1317 0.2843 0.2898 0.2335 0 0	0 0 -16.2665 0.0710 0 0 1.0000 0 21.3000 4.5000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 1.0000\\ 8.0000\\ 7.0000\\ 6.0000\\ 21.3000\\ 4.5000\end{array}$
2005	LV	2704	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	2704 2704 2704 2704 2704 2686 2704 2704 2704 2704 2704	0 0 0 0 18 0 0 0 0 0 0	0.8089 0.5195 -0.6736 78.6490 0.5020 0.4826 0.3015 0.7461 1.6262 10.9000 8.9000	$\begin{array}{c} 0.2019\\ 0.2566\\ 4.5409\\ 37.5947\\ 0.2568\\ 0.2566\\ 0.2357\\ 0.4546\\ 0.5490\\ 0.4794\\ 0\\ 0\\ 0\end{array}$	0 0 -16.2665 0.0710 0 0 0 1.0000 0 10.9000 8.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 6.0000 5.0000 10.9000 8.9000
2006	LV	3438	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	3438 3438 3438 3438 3438 3438 3417 3438 3438 3438 3438 3438 3438	0 0 0 0 21 0 0 0 0 0	0.7396 0.5180 -0.7024 79.0465 0.5185 0.4850 0.3494 0.7660 2.6926 1.7491 12.4000 6.8000	0.1993 0.2270 4.0261 32.8692 0.2270 0.2270 0.2270 0.2165 0.4175 0.4952 0.4339 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 12.4000 6.8000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 6.0000 5.0000 12.4000 6.8000
2005	LT	2527	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	2527 2527 2527 2527 2527 2527 2455 2527 2527	0 0 0 0 72 0 0 0 0 0 0	0.7906 0.5165 -0.8742 75.0490 0.3968 0.3962 0.3853 0.9105 2.5087 1.6172 11.9000 8.3000	0.2636 0.3237 5.5828 46.8519 0.3169 0.3168 0.3162 0.6330 0.6445 0.5482 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 11.9000 8.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 9.0000 9.0000 5.0000 11.9000 8.3000
2006	LT	3713	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	3713 3713 3713 3713 3713 3613 3713 3713	0 0 0 0 100 0 0 0 0 0 0 0	0.7549 0.5173 -0.6671 74.1260 0.4288 0.3475 0.4118 0.9010 2.5249 1.6651 13.1000 5.6000	0.2280 0.2649 4.5500 37.1506 0.2623 0.2524 0.2618 0.5201 0.5195 0.4300 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 13.1000 5.6000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 9.0000 9.0000 5.0000 13.1000 5.6000
2005	LU	4322	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH</pre>	4322 4322 4322 4322 4322 4322 4322 4322	0 0 0 0 0 112	0.0106 0.4941 -0.5958 70.9201 0.3231 0.3213 0.3544	0.0191 0.0930 1.5620 12.6713 0.0870 0.0868 0.0895	0 0 -16.2665 0.0710 0 0 0	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000

			nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	4322 4322 4322 4322 4322 4322	0 0 0 0	0.9074 2.3891 1.5991 57.1000 4.5000	0.2003 0.1848 0.1397 0 0	0 1.0000 0 57.1000 4.5000	7.0000 7.0000 5.0000 57.1000 4.5000
2006	LU	4601	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	4601 4601 4601 4601 4601 4601 4601 4601	0 0 0 0 107 0 0 0 0 0	0.0192 0.4979 -0.8068 73.1864 0.3346 0.3191 0.3524 0.9019 2.3624 1.5974 63.1000 4.7000	0.0251 0.0917 1.5614 12.9800 0.0865 0.0865 0.0882 0.1980 0.1804 0.1353 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 63.1000 4.7000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 6.0000\\ 6.0000\\ 6.0000\\ 5.0000\\ 63.1000\\ 4.7000\end{array}$
2005	ΗU	4353	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	4353 4353 4353 4353 4353 4353 4353 4353	0 0 0 0 0 3 0 0 0 0 0 0 0	0.7286 0.5111 -0.7273 86.0190 0.5287 0.4448 0.2556 0.8801 2.6492 1.5751 14.2595 7.3201	0.3793 0.4264 7.8870 62.2732 0.4258 0.4239 0.3721 0.9083 0.8968 0.7810 4.9600 1.4527	0 0 -16.2665 0.0710 0 0 0 1.0000 0 9.3000 5.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 7.0000 9.0000 5.0000 23.2000 9.2000
2006	HU	6609	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	6609 6609 6609 6609 6609 6609 6609 6609	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	0.6904 0.5155 -0.7190 85.6337 0.4654 0.3874 0.2955 0.8703 2.6738 1.5826 15.0909 7.5619	0.3197 0.3456 6.3795 50.2613 0.3449 0.3369 0.3155 0.7353 0.7254 0.6395 4.4141 1.4049	0 0 -16.2665 0.0710 0 0 0 1.0000 0 9.7000 5.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 7.0000 5.0000 24.9000 9.9000
2005	NL	8574	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	8574 8574 8574 8574 8574 8206 8574 8574 8407 8574 8574	0 0 0 368 0 167 0	0.0581 0.4921 -0.4148 74.2244 0.2497 0.1563 0.4600 0.9032 2.1546 1.4760 29.4000 4.7000	0.1868 0.3992 6.8706 56.6392 0.3456 0.2899 0.3967 0.8799 0.6870 0.6251 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 29.4000 4.7000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 0.0000 7.0000 7.0000 5.0000 29.4000 4.7000
2006	NL	9287	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	9287 9287 9287 9287 9287 8949 9287 9287 9287 9279 9287 9287	0 0 0 338 0 0 8 0 0	0.0318 0.4975 -0.2668 73.4840 0.2542 0.1644 0.4702 0.9265 2.1678 1.6542 30.9000 3.9000	0.1310 0.3732 6.3946 53.0152 0.3249 0.2766 0.3716 0.8325 0.6369 0.5648 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 30.9000 3.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 7.0000 7.0000 5.0000 30.9000 3.9000
2005	AT	3882	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	3882 3882 3882 3882 3882 3882 3882 3882	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	0.0522 0.5056 -0.1534 70.8171 0.2931 0.3432 0.3181 0.8440 2.3847 1.6527 28.0319 5.1660	0.1828 0.4107 6.9123 57.3363 0.3739 0.3900 0.3826 0.8552 0.8766 0.7373 1.8223 1.0814	0 0 -16.2665 0.0710 0 0 0 1.0000 0 23.9000 3.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 8.0000 7.0000 29.6000 6.7000
2006	AT	5469	pov_indicator woman age_centered age_centered2	5469 5469 5469 5469	0 0 0 0	0.0555 0.5080 -0.3177 70.2327	0.1577 0.3443 5.7667 48.1919	0 0 -16.2665 0.0710	1.0000 1.0000 14.7335 264.6

			chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts	5469 5469 5469 5469 5469 5469 5469 5469 5469		$\begin{array}{c} 0.2754 \\ 0.3202 \\ 0.3149 \\ 0.8224 \\ 2.3748 \\ 1.6395 \\ 29.4096 \\ 4.7459 \end{array}$	0.3076 0.3213 0.3198 0.7036 0.7198 0.6173 1.4577 0.9453	0 0 1.0000 25.3000 3.3000	1.0000 1.0000 6.0000 8.0000 6.0000 30.8000 6.3000
2005	ΡL	10627	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	10627 10627 10627 10627 10627 10561 10627 10564 10627 10564 10627	0 0 0 66 0 63 0 0	0.7381 0.5039 -0.6632 85.8863 0.4761 0.2527 0.2646 0.8989 2.7400 1.4003 11.5603 17.7915	0.4729 0.5378 9.9424 79.9035 0.5372 0.4674 0.4748 1.1369 1.3407 1.0024 2.5551 2.0488	0 0 -16.2665 0.0710 0 0 0 1.0000 0 8.2000 15.7000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 1.0000\\ 1.0000\\ 9.0000\\ 5.0000\\ 15.7000\\ 21.4000\end{array}$
2006	PL	14644	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	14644 14644 14644 14644 12928 14644 14644 14644 14644 13001 14644 14644	0 0 0 1716 0 1643 0 0	0.7258 0.5037 -0.6473 85.1495 0.4608 0.3260 0.2834 0.9003 2.9673 1.5990 12.3506 13.9915	0.4102 0.4598 8.4644 67.2797 0.4594 0.4310 0.4143 0.9688 1.1522 0.8760 2.4094 1.0485	0 0 -16.2665 0.0710 0 0 0 1.0000 0 8.7000 12.7000	1.0000 14.7335 264.6 1.0000 1.0000 1.0000 1.0000 9.0000 6.0000 17.0000 16.4000
2005	ΡT	2671	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	2671 2671 2671 2671 2671 2198 2671 2671 2671 2671 2671	$\begin{array}{c} 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 473 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{array}$	0.4168 0.5072 -0.6899 79.3436 0.4695 0.4113 0.2303 0.8315 2.8063 1.8021 17.3000 7.6000	0.5441 0.5517 9.8001 79.7975 0.5507 0.5430 0.4696 0.9949 1.1722 1.0142 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 1.0000 0 17.3000 7.6000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 9.0000\\ 8.0000\\ 7.0000\\ 7.0000\\ 7.0000\\ 7.6000\end{array}$
2006	ΡΤ	3662	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	3662 3662 3662 3662 3662 3049 3662 3662 3662 3662 3662 3662 3662	0 0 0 0 613 0 0 0 0 0 0	0.4019 0.5112 -0.3565 79.3109 0.4380 0.3936 0.2419 0.7832 2.8086 1.8287 18.0000 7.7000	0.4575 0.4665 8.3039 66.5722 0.4630 0.4559 0.4037 0.8223 1.0182 0.8564 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 18.0000 7.7000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 9.0000\\ 9.0000\\ 9.0000\\ 7.0000\\ 18.0000\\ 7.7000\end{array}$
2005	SI	8671	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	8671 8667 8671 8671 8671 8671 8671 8671	0 4 0 0 200 0 200 0 0 13 0 0	0.1266 0.4899 -0.6833 79.6898 0.2501 0.2177 0.2149 0.7566 2.9380 1.7535 19.6000 6.5000	0.0915 0.1376 2.4495 19.9780 0.1192 0.1136 0.1128 0.2591 0.3047 0.2551 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 19.6000 6.5000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 9.0000 9.0000 9.0000 6.0000 19.6000 6.5000
2006	SI	11271	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	11271 11263 11271 11271 11271 11271 11271 11271 11262 11271 11271	0 8 0 0 82 0 9 0 9 0	0.1306 0.4929 -0.7373 78.4043 0.2915 0.2126 0.2549 0.7634 2.9358 1.7851 20.7000 6.0000	0.0808 0.1200 2.1167 17.6044 0.1090 0.0981 0.1046 0.2254 0.2660 0.2174 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 20.7000 6.0000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 9.0000 9.0000 9.0000 6.0000 20.7000 6.0000

2005	SK	3447	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	3447 3447 3447 3447 3447 3432 3447 3447	0 0 0 0 15 0 0 0 0 0 0	0.7661 0.5198 0.1174 81.3319 0.4441 0.3993 0.3085 0.8904 3.0942 1.9700 13.5000 16.3000	0.3005 0.3547 6.4023 53.0850 0.3528 0.3477 0.3279 0.7283 0.8632 0.7508 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 1.0000 0 13.5000 16.3000	$\begin{array}{c} 1.0000\\ 1.0000\\ 14.7335\\ 264.6\\ 1.0000\\ 1.0000\\ 1.0000\\ 6.0000\\ 8.0000\\ 8.0000\\ 13.5000\\ 16.3000\end{array}$
2006	SK	4898	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	4898 4898 4898 4898 4898 4898 4898 4898	0 0 0 28 0 0 0 0 0 0	0.7157 0.5217 0.0502 82.3628 0.4327 0.4442 0.3248 0.8442 3.1409 2.0315 15.0000 13.4000	0.2680 0.2968 5.3919 44.6004 0.2952 0.2952 0.2783 0.5925 0.7288 0.6170 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 15.0000 13.4000	1.0000 14.7335 264.6 1.0000 1.0000 8.0000 8.0000 8.0000 8.0000 15.0000 13.4000
2005	FI	4021	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	4021 4021 4021 4021 4021 4021 4021 4021	0 0 0 94 0 0 3 0 0	0.0529 0.4953 -0.3310 80.6732 0.3076 0.2930 0.5226 0.8869 2.1009 1.4707 25.6182 8.5095	0.1404 0.3137 5.6323 46.9628 0.2896 0.2856 0.3133 0.7442 0.5226 0.4894 2.5938 1.1694	0 0 -16.2665 0.0710 0 0 0 1.0000 0 19.1000 6.9000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 7.0000 7.0000 5.0000 29.6000 11.7000
2006	FI	5528	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	5528 5528 5528 5528 5528 5528 5528 5528	0 0 0 0 124 0 0 0 0 0	0.0622 0.4953 -0.3237 80.8606 0.2901 0.2894 0.5179 0.8898 2.0887 1.4683 27.1434 7.7955	0.1290 0.2669 4.7976 39.6984 0.2423 0.2421 0.2668 0.6325 0.4440 0.3940 2.3172 0.9975	0 0 -16.2665 0.0710 0 0 0 1.0000 0 20.2000 6.3000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 1.0000 12.0000 8.0000 4.0000 31.3000 11.3000
2005	SE	3541	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	3541 3541 3541 3541 3541 3541 3541 3541	0 0 0 0 127 0 0 386 0 0	0.0452 0.5016 -0.3763 78.1345 0.2827 0.1482 0.4394 0.9789 2.1022 1.5457 27.1000 7.5000	0.1750 0.4211 7.4383 60.4577 0.3793 0.2992 0.4189 0.9351 0.6935 0.5973 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 0 27.1000 7.5000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 8.0000 7.0000 5.0000 27.1000 7.5000
2006	SE	4751	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH GDPhabnuts unempratenuts</pre>	4751 4751 4751 4751 4751 4534 4751 4751 4210 4751 4751	$\begin{array}{c} 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 217 \\ 0 \\ 541 \\ 0 \\ 0 \\ 0 \end{array}$	0.0527 0.5011 -0.4989 76.0001 0.4920 0.1593 0.4728 1.0203 2.0672 1.5782 28.7000 7.1000	0.1577 0.3531 6.1470 49.7601 0.3211 0.2585 0.3530 0.7818 0.5636 0.4843 0 0 0	0 0 -16.2665 0.0710 0 0 0 1.0000 28.7000 7.1000	1.0000 1.0000 14.7335 264.6 1.0000 1.0000 6.0000 7.0000 28.7000 7.1000
2005	UK	6224	<pre>pov_indicator woman age_centered age_centered2 chronicdiseaseHH activityhamperedHH uppereducHH nbchildren nbadultsHH nbemployedHH</pre>	6224 6224 6224 6224 6224 6224 4949 6224 6224	0 0 0 1275 0 0 1458	$\begin{array}{c} 0.0739\\ 0.5260\\ -1.9273\\ 67.7453\\ 0.4031\\ 0.2293\\ 0.5177\\ 1.2542\\ 2.1928\\ 1.5478\end{array}$	0.5272 1.0064 16.1286 142.1 0.9887 0.8473 1.0064 2.3990 1.7624 1.7058	0 0 -16.2665 0.0710 0 0 0 0 1.0000 0	1.0000 1.0000 13.7335 264.6 1.0000 1.0000 1.0000 6.0000 7.0000 6.0000

			GDPhabnuts unempratenuts	6224 6224	0	27.4000 4.8000	0	27.4000 4.8000	27.4000 4.8000
			unempracentucs	0224	0	4.0000	0	4.0000	4.0000
2006	UK	7594	pov_indicator	7594	0	0.0837	0.4988	0	1.0000
			woman	7594	0	0.5363	0.8983	0	1.0000
			age_centered	7594	0	-1.5980	14.5325	-16.2665 0.0710	13.7335 264.6
			age_centered2 chronicdiseaseHH	7594 7594	0	67.6428 0.4049	128.3 0.8842	0.0710	264.6
			activityhamperedHH	7594	0	0.2199	0.7461	0	1.0000
			uppereducHH	5998	1596	0.5080	0.8985	0	1.0000
			nbchildren	7594	0	1.2305	2.1147	0	6.0000
			nbadultsHH	7594	0	2.1941	1.5651	1.0000	8.0000
			nbemployedHH	6118	1476	1.5570	1.4788	0	5.0000
			GDPhabnuts	7594	0	28.4000	0	28.4000	28.4000
			unempratenuts	7594	0	5.4000	0	5.4000	5.4000
2005	IS	1583	pov_indicator	1583	0	0.0333	0.0418	0	1.0000
			woman	1583	0	0.5066	0.1164	0	1.0000
			age_centered	1583 1583	0	-1.0735 76.5700	2.0227 16.9520	-16.2665 0.0710	14.7335 264.6
			age_centered2 chronicdiseaseHH	1583	0	0.2143	0.0956	0.0710	1.0000
			activityhamperedHH	1583	0	0.1521	0.0836	0	1.0000
			uppereducHH	1359	224	0.3835	0.1150	0	1.0000
			nbchildren	1583	0	1.1989	0.2545	0	5.0000
			nbadultsHH	1583	0	2.4249	0.2303	1.0000	7.0000
			nbemployedHH	1318	265	1.8715	0.1947	0	6.0000
			GDPhabnuts unempratenuts	1583 1583	0	29.3000 2.5000	0	29.3000 2.5000	29.3000 2.5000
			unempracentics	1000	0	2.3000	0	2.3000	2.5000
2006	IS	2240	pov_indicator	2240	0	0.0322	0.0348	0	1.0000
			woman age centered	2240 2240	0	0.4993 -0.9177	0.0986 1.7342	0 -16.2665	1.0000 14.7335
			age_centered2	2240	0	78.2001	14.4429	0.0710	264.6
			chronicdiseaseHH	2240	Ő	0.2180	0.0814	010110	1.0000
			activityhamperedHH	2240	0	0.1579	0.0719	0	1.0000
			uppereducHH	2021	219	0.4147	0.0983	0	1.0000
			nbchildren	2240	0	1.1305	0.2161	0	5.0000
			nbadultsHH	2240	0	2.3563	0.1875	1.0000	7.0000
			nbemployedHH GDPhabnuts	2054 2240	186 0	1.7410 29.3000	0.1499 0	0 29.3000	5.0000 29.3000
			unempratenuts	2240	0	2.8000	0	2.8000	2.8000
2005	NO	4088	pov indicator	4088	0	0.0326	0.1001	0	1.0000
			woman	4088	0	0.4919	0.2816	0	1.0000
			age_centered	4088	0	-0.6817	4.8555	-16.2665	14.7335
			age_centered2	4088	0	74.7612	40.0863	0.0710	264.6
			chronicdiseaseHH	4088	0	0.2545	0.2454 0.2070	0	1.0000
			activityhamperedHH uppereducHH	4088 3887	201	0.1610 0.3970	0.2070	0	1.0000 1.0000
			nbchildren	4088	0	1.0059	0.6332	õ	8.0000
			nbadultsHH	4088	0	2.0123	0.4476	1.0000	8.0000
			nbemployedHH	3996	92	1.4737	0.4073	0	4.0000
			GDPhabnuts	4088	0	39.6000	0	39.6000	39.6000
			unempratenuts	4088	0	4.4000	0	4.4000	4.4000
2006	NO	3926	pov_indicator	3926	0 0	0.0312	0.0978	0 0	1.0000
			woman age centered	3926 3926	0	0.4909 -0.5332	0.2814 4.7661	-16.2665	1.0000 14.7335
			age_centered2	3926	0	71.9978	39.3175	0.0710	264.6
			chronicdiseaseHH	3926	0	0.2769	0.2518	0	1.0000
			activityhamperedHH	3926	0	0.1713	0.2120	0	1.0000
			uppereducHH	3583	343	0.4123	0.2776	0	1.0000
			nbchildren	3926	0	0.9872	0.6287	0	8.0000
			nbadultsHH nbemployedHH	3926 3732	0 194	2.0046 1.5183	0.4367 0.4010	1.0000	8.0000 5.0000
			GDPhabnuts	3926	194	43.4000	0.4010	43.4000	43.4000
			unempratenuts	3926	0	3.4000	0	3.4000	3.4000
			-						

#### Appendix 6. Empty model (weighted)

			Response 1	Profile		
		Ordered Value	pov_indi	cator F	Total requency	
		1 2	1 0		76252 222018	
Th	e GLIMMIX pro	cedure is mode	ling the p	robability	' that pov_in	dicator='1'.
		G-side C	Dimens: ov. Paramet		1	
		Columns	ov. Paramet		2 1	
		Columns	in Z per Si	ubject	1	
		Subjects	(Blocks 1)	nV)	93	
		Max Obs	per Subject	t	19942	
			imization 3			
		Optimization Parameters in	Technique Optimizat:	New ion 2	ton-Raphson	
		Lower Boundar	ies	2		
		Upper Boundar		1		
		Fixed Effects			filed	
		Residual Vari Starting From			filed 1 estimates	
/ 6 iteratio		-				
	Con	vergence crite	rion (PCON	V=1.11022E	1-8) satisfie	d.
			Fit Stat:			
		-2 Res Log			1738618 294644.1	
		Generalized Gener. Chi-			0.99	
			nce Paramet		+	
		Covaria.	nce Parame	ter Bstima		
						ndard
		rm Subject			Sta	
	Interc	rm Subject ept region			Sta	
	Interc AR(1)	rm Subject ept region ID_uniq al	ue_UE(regio	Est 1 on) 0	Sta 	
	Interc AR(1) Residu	al		Est 1 on) C C	Sta .7720 0 .4571 0.0 .9878 0.0	Error .2639 02340 02734
	Interc AR(1) Residu Asymptoti	al c Covariance M	atrix of Co	Est 1 on) 0 0 ovariance	Sta imate .7720 0 .4571 0.0 .9878 0.0 Parameter Es	Error .2639 02340 02734 timates
	Interc AR(1) Residu Asymptoti Cov Parm	al c Covariance M Subject	atrix of Co	Est 1 on) C 0 variance CovP1	Sta imate .7720 0 .4571 0.0 .9878 0.0 Parameter Es CovP2	Error .2639 02340 02734 timates CovP3
	Interc AR(1) Residu Asymptoti Cov Parm Intercept AR(1)	al c Covariance M Subject	atrix of Co	Est 1 on) C 0 variance CovP1	Sta imate .7720 0 .4571 0.0 .9878 0.0 Parameter Es CovP2	Error .2639 02340 02734 timates CovP3
	Interc AR(1) Residu Asymptoti Cov Parm Intercept	al c Covariance M Subject	atrix of Co	Est 1 on) C 0 variance CovP1	Sta imate .7720 0 .4571 0.0 .9878 0.0 Parameter Es CovP2	Error .2639 02340 02734 timates CovP3
	Interc AR(1) Residu Asymptoti Cov Parm Intercept AR(1) Residual Asymptoti	al c Covariance M Subject region ID_unique_UE( c Correlation 1	atrix of Co region)	Est 1 0n) C 0 0 0 0 0 0 0 0 0 0 0 0 0	Sta imate .7720 0 .4571 0.0 .9878 0.0 Parameter Es CovP2 -2.32E-7 5.477E-6 2.254E-6 Parameter E	Error .2639 02340 02734 timates CovP3 -2.12E-7 2.254E-6 7.473E-6 stimates
	Interc AR(1) Residu Asymptoti Cov Parm Intercept AR(1) Residual Asymptoti Cov Parm	al c Covariance M Subject region ID_unique_UE( c Correlation M Subject	atrix of Co region)	Est 1 0n) C 0 0 0 0 0 0 0 0 0 0 0 0 0	Sta imate .7720 0 .4571 0.0 .9878 0.0 Parameter Es CovP2 -2.32E-7 5.477E-6 2.254E-6 Parameter E	Error .2639 02340 02734 timates CovP3 -2.12E-7 2.254E-6 7.473E-6 stimates
	Interc AR(1) Residu Asymptoti Cov Parm Intercept AR(1) Residual Asymptoti Cov Parm Intercept	al c Covariance M Subject region ID_unique_UE( c Correlation 1 Subject region	atrix of Co region) Matrix of (	Est 1 0 0 0 0 0 0 0 0 0 0 0 0 0	Sta 	Error .2639 02340 02734 timates CovP3 -2.12E-7 2.254E-6 7.473E-6 stimates CovP3 -0.00029
	Interc AR(1) Residu Asymptoti Cov Parm Intercept AR(1) Residual Asymptoti Cov Parm	al c Covariance M Subject region ID_unique_UE( c Correlation M Subject	atrix of Co region) Matrix of (	Est 1 0 0 0 0 0 0 0 0 0 0 0 0 0	Sta 	Error .2639 02340 02734 timates CovP3 -2.12E-7 2.254E-6 7.473E-6 stimates CovP3 -0.00029
	Interc AR(1) Residu Asymptoti Cov Parm Intercept AR(1) Residual Asymptoti Cov Parm Intercept AR(1)	al c Covariance M Subject region ID_unique_UE( c Correlation M Subject region ID_unique_UE(	atrix of Co region) Matrix of ( region) ions for F:	Est 1 000) 0 000000000000000000000000000000	Sta imate 	Error .2639 02340 02734 timates CovP3 -2.12E-7 2.254E-6 7.473E-6 stimates CovP3 -0.00029
	Interc AR(1) Residu Asymptoti Cov Parm Intercept AR(1) Residual Asymptoti Cov Parm Intercept AR(1)	al c Covariance M Subject region ID_unique_UE( c Correlation M Subject region ID_unique_UE( Solut	atrix of Co region) Matrix of ( region) ions for F: Standard	Est 1 000) 0 000000000000000000000000000000	Sta 	Error .2639 02340 02734 timates 

The empty model contains only a random intercept. The dependent variable is thus explained by the overall mean, a random term at group level and a random term at individual level. The empty model allows to find out the relative parts of the variance between groups (regions here) and the variance within groups. It is thus possible to calculate the intraclass correlation coefficient (ICC) (here: 55.1% = 1.772/3.2169, i.e. the intercept variance divided by the total variance).



B.P. 48 L-4501 Differdange Tél.: +352 58.58.55-801 www.ceps.lu